

Lecture 6: More about Difference Equations

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Disclaimer: Zhikun is fully responsible for the errors and typos appeared in the notes.

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6.1 First Order Linear Difference Equations

$$\underbrace{y_t}_{\text{Endog.}} = \lambda y_{t-1} + b \underbrace{x_t}_{\text{Exog.}} + a, \quad a, b, \lambda = \text{parameters, } \lambda \neq 1 \quad (6.1)$$

Note:

$\lambda = 1$ unit root, non-stationary, random walk

Sargent 1979, CH.9

$$y_t - \lambda y_{t-1} = b x_t + a \quad (6.2)$$

$$(1 - \lambda L)y_t = b x_t + a \quad (6.3)$$

$$y_t = \frac{1}{1 - \lambda L}(b x_t + a) + c \lambda^t \quad (6.4)$$

where $c \lambda^t$ is the “summation constant”, similar to the constant in indefinite integral. We can check that

$$(1 - \lambda L)c \lambda^t = c \lambda^t - \lambda L(c \lambda^t) = c \lambda^t - \lambda c \lambda^{t-1} = 0$$

Note

$$L(x_t \pm y_t) = x_{t-1} \pm y_{t-1} \quad (6.5)$$

$$L(a x_t) = a L x_t = a x_{t-1} \quad (6.6)$$

$$y_t = \left(\frac{1}{1 - \lambda L}\right) b x_t + \left(\frac{1}{1 - \lambda L}\right) a + c \lambda^t \quad (6.7)$$

Assuptions

$|\lambda| < 1 \implies L a = a$.

Then

$$\left(\frac{1}{1 - \lambda L}\right) a = \left(\frac{1}{1 - \lambda}\right) a \quad (6.8)$$

This is because

$$\left(\frac{1}{1 - \lambda L}\right) a = \left(\sum_{i=0}^{\infty} \lambda^i L^i\right) a = \sum_{i=0}^{\infty} (\lambda^i L^i a) = \sum_{i=0}^{\infty} (\lambda^i a) = a \sum_{i=0}^{\infty} \lambda^i = a \left(\frac{1}{1 - \lambda}\right) \quad (6.9)$$

¹Visit <http://www.luzk.net/misc> for updates.

Thus,

$$\begin{aligned} y_t &= \left(\frac{1}{1-\lambda L}\right)bx_t + \left(\frac{1}{1-\lambda L}\right)a + c\lambda^t \\ &= b \sum_{i=0}^{\infty} \lambda^i x_{t-i} + \frac{a}{1-\lambda} + c\lambda^t \end{aligned} \quad (6.10)$$

Recall

If $|\lambda| < 1$, then $\left(\frac{1}{1-\lambda L}\right)x_t = \sum_{i=0}^{\infty} \lambda^i x_{t-i}$

If $|\lambda| > 1$, then $\left(\frac{1}{1-\lambda L}\right)x_t = -\sum_{i=1}^{\infty} \left(\frac{1}{\lambda}\right)^i x_{t+i}$

Note: If $|\lambda| > 1$, then we would use forward-looking solution.

Example: Cagan (1956)

$$\begin{cases} m_t &= \ln M_t \\ p_t &= \ln P_t \\ p_{t+1}^e &= \ln P_{t+1}^e \end{cases} \quad (6.11)$$

Money Market: Money demand

$$\underbrace{m_t - p_t}_{\ln\left(\frac{M_t}{P_t}\right)} = \alpha \left(\underbrace{p_{t+1}^e - p_t}_{\text{expeced inflation}} \right), \quad \alpha < 0 \quad (6.12)$$

this is a(n) simplification/approximation under hyperinflation:

$$\left(\frac{M}{P}\right)^d = \frac{M}{P}(i, y) = \frac{M}{P}(r + \pi^e, y) \quad (6.13)$$

When inflation is high, r and y can be treated as fixed.

Assumption

$$y = \bar{y}, r = \bar{r}$$

$$P_{t+1}^e = (1 - \gamma)P_t + \gamma P_{t-1} \implies \pi_{t+1}^e = \Gamma \pi_t, \quad \Gamma < 0 \quad (6.14)$$

which is an example of extrapolative expectation.

$$p_{t+1}^e = (1 + \Gamma)p_t - \Gamma p_{t-1} \quad (6.15)$$

$$m_t - p_t = \alpha \Gamma (p_t - p_{t-1}) \quad (6.16)$$

$$\implies p_t - \underbrace{\left(\frac{\alpha \Gamma}{1 + \alpha \Gamma}\right)}_{\lambda} p_{t-1} = \frac{1}{1 + \alpha \Gamma} m_t \quad (6.17)$$

$$p_t - \lambda p_{t-1} = \frac{1}{1 + \alpha \Gamma} m_t \quad (6.18)$$

$$p_t = \frac{1}{1 + \alpha \Gamma} \left(\frac{1}{1 - \lambda L}\right) m_t + c\lambda^t \quad (6.19)$$

Since $|\lambda| = \left|\frac{\alpha \Gamma}{1 + \alpha \Gamma}\right| < 1$, \implies

$$p_t = \frac{1}{1 + \alpha \Gamma} \sum_{i=0}^{\infty} \lambda^i m_{t-i} + c\lambda^t = \frac{1}{1 + \alpha \Gamma} \sum_{i=0}^{\infty} \left(\frac{\alpha \Gamma}{1 + \alpha \Gamma}\right)^i m_{t-i} + c\left(\frac{\alpha \Gamma}{1 + \alpha \Gamma}\right)^t \quad (6.20)$$

Example: Perfect foresight

$$p_{t+1}^e - p_t = p_{t+1} - p_t \quad (6.21)$$

$$\pi_{t+1}^e = \pi_{t+1} \quad (6.22)$$

Then money demand is

$$m_t - p_t = \alpha(p_{t+1} - p_t) \quad (6.23)$$

$$\implies p_{t+1} - \frac{\alpha - 1}{\alpha} p_t = \frac{m_t}{\alpha} \quad (6.24)$$

$$\implies (1 - \lambda L)p_{t+1} = \frac{m_t}{\alpha} \quad (6.25)$$

Here, $|\lambda| = \left| \frac{\alpha - 1}{\alpha} \right| > 1$. Hence use backward-looking solution:

$$\begin{aligned} p_{t+1} &= -\frac{1}{\alpha} \left(\frac{1}{1 - \lambda L} \right) m_t + c\lambda^t \\ &= -\frac{1}{\alpha} \sum_{i=1}^{\infty} \left(\frac{1}{\lambda} \right)^i m_{t+i} + c\lambda^t \\ &= -\frac{1}{\alpha} \sum_{i=1}^{\infty} \left(\frac{1}{\left(\frac{\alpha - 1}{\alpha} \right)} \right)^i m_{t+i} + c \left(\frac{\alpha - 1}{\alpha} \right)^t \\ &= \frac{1 - \alpha}{\alpha^2} \sum_{i=1}^{\infty} \left(\frac{\alpha}{\alpha - 1} \right)^{i+1} m_{t+i} + c \left(\frac{\alpha - 1}{\alpha} \right)^t \\ p_t &= \frac{1 - \alpha}{\alpha^2} \sum_{i=1}^{\infty} \left(\frac{\alpha}{\alpha - 1} \right)^{i+1} m_{t+i-1} + c \left(\frac{\alpha - 1}{\alpha} \right)^{t-1} \end{aligned} \quad (6.26)$$

6.2 Dynamic Discrete Time Infinite Horizon Model (Ramsey)

$$\max \sum_{t=1}^{\infty} \beta^t u(c_t), \quad 1 > \beta > 0 \quad (6.27)$$

$$\text{s.t. } c_t + B_t = (1 + \rho_{t-1})B_{t-1} + y_t \quad (6.28)$$

$$\mathcal{L} = \sum_{t=1}^{\infty} \beta^t u(c_t) - \sum_{t=1}^{\infty} \lambda_t [(1 + \rho_{t-1})B_{t-1} + y_t - c_t - B_t] \quad (6.29)$$

Choice: $\{c_t, B_t\}_{t=1}^{\infty}$

FONC:

$$[c_t] \quad \beta^t u'(c_t) - \lambda_t = 0 \quad (6.30)$$

$$[B_t] \quad -\lambda_t + \lambda_{t+1}(1 + \rho_t) = 0 \quad (6.31)$$

\implies

$$1 + \rho_t = \frac{u'(c_t)}{\beta u'(c_{t+1})} \quad (6.32)$$

which says “objective rate of substitution” = “subjective rate of substitution”.

Assumption: $\rho_t = \rho, \forall t$

$$(1 - (1 + \rho)L)B_t = y_t - c_t \quad (6.33)$$

$$B_t = \frac{1}{1 - (1 + \rho)L}(y_t - c_t) + d(1 + \rho)^t \quad (6.34)$$

$$= -\sum_{s=1}^{\infty} \left(\frac{1}{1 + \rho}\right)^s (y_{t+s} - c_{t+s}) + d(1 + \rho)^t \quad (6.35)$$

$$= -\sum_{s=1}^{\infty} \left(\frac{1}{1 + \rho}\right)^s (y_{t+s} - c_{t+s}) \quad (d = 0 \text{ because of NPGC}) \quad (6.36)$$

$$\sum_{s=1}^{\infty} \left(\frac{1}{1 + \rho}\right)^s c_{t+s} = B_t + \sum_{s=1}^{\infty} \left(\frac{1}{1 + \rho}\right)^s y_{t+s} \quad (6.37)$$

which is the life-time budget constraint.

NPGC:

$$\lim_{t \rightarrow \infty} d(1 + \rho)^t = 0 \implies d = 0 \quad (6.38)$$

Assumption $y_t = y, u(c) = \ln(c) \implies$

$$1 + \rho = \frac{1}{\beta} \frac{c_{t+1}}{c_t} \iff (1 + \rho)\beta c_t = c_{t+1} \quad (6.39)$$

$$c_{t+1} = (1 + \rho)\beta c_t \quad (6.40)$$

$$c_{t+2} = ((1 + \rho)\beta)^2 c_t \quad (6.41)$$

$$c_{t+3} = ((1 + \rho)\beta)^3 c_t \quad (6.42)$$

$$\vdots \quad (6.43)$$

$$c_{t+s} = ((1 + \rho)\beta)^s c_t \quad (6.44)$$

\implies

$$B_t + \sum_{s=1}^{\infty} \left(\frac{1}{1 + \rho}\right)^s y = \sum_{s=1}^{\infty} \left(\frac{1}{1 + \rho}\right)^s (1 + \rho)^s \beta^s c_t \quad (6.45)$$

$$B_t + y \sum_{s=1}^{\infty} \left(\frac{1}{1 + \rho}\right)^s = c_t \sum_{s=1}^{\infty} \left(\frac{1}{1 + \rho}\right)^s (1 + \rho)^s \beta^s \quad (6.46)$$

$$B_t + \frac{1}{\rho} y = \frac{\beta}{1 - \beta} c_t \quad (6.47)$$

$$c_t = \frac{1 - \beta}{\beta \rho} (y + \rho B_t) \quad (6.48)$$

6.3 Expectations

How to get information about expectation?

1. Survey
2. Observe behaviour
3. Assume some mechanism
 - (a) Static expectation $p_{t+1}^e = p_t$
Example: The cobweb model
 - (b) Extrapolative expectations
Example:

$$p_{t+1}^e = \Gamma p_t + (1 - \Gamma)p_{t-1} = \Gamma(p_t - p_{t-1}) + p_{t-1} \quad (6.49)$$

Example:

$$\begin{aligned} C_{t+1} &= \alpha + \beta y_{t+1}^e \\ &= \alpha + \beta \Gamma y_t + \beta(1 - \Gamma)y_{t-1} + \epsilon_t \end{aligned} \quad (6.50)$$

- (c) Adaptive expectations

Adaptive Expectations

$$p_{t+1}^e = p_t^e + \theta \underbrace{(p_t - p_t^e)}_{\text{forecast error}}, \quad 0 < \theta < 1 \quad (6.51)$$

$$p_{t+1}^e = \theta \sum_{s=0}^{\infty} (1 - \theta)^s p_{t-s} + c(1 - \theta)^t \quad (6.52)$$

Rational Expectations

- 1) $P_{t+1}^e = \mathbb{E}[P_{t+1} | \Omega_t]$
- 2) $P_{t+1}^e - P_{t+1} = \epsilon_{t+1}$, where $\mathbb{E}(\epsilon_{t+1}) = 0$, $\text{cov}(\epsilon_t, \epsilon_{t\pm 1}) = 0$.

Perfect Foresight is an extreme case of RE where $\epsilon_{t+1} = 0 \forall t$.

6.4 Stochastic Difference Equation

$$y_t = a \underbrace{\mathbb{E}[y_{t+1} | I_t]}_{\mathbb{E}_t y_{t+1}} + c x_t \quad (6.53)$$

Example:

$$p_t = \left(\frac{\alpha}{1 + \alpha}\right) \mathbb{E}[p_{t+1} | I_t] + \left(\frac{1}{1 + \alpha}\right) m_t \quad (6.54)$$

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Law of Iterated Expectations

$$\mathbb{E}[\mathbb{E}[X | I_{t+1}] | I_t] = \mathbb{E}[X | I_t] \quad (6.55)$$

$$y_t = a\mathbb{E}_t y_{t+1} + cx_t \quad (6.56)$$

$$y_{t+1} = a\mathbb{E}_{t+1} y_{t+2} + cx_{t+1} \quad (6.57)$$

$$\mathbb{E}_t y_{t+1} = a\mathbb{E}_t \mathbb{E}_{t+1} y_{t+2} + c\mathbb{E}_t x_{t+1} = a\mathbb{E}_t y_{t+2} + c\mathbb{E}_t x_{t+1} \quad (6.58)$$

$$\implies y_t = a^2 \mathbb{E}_t y_{t+2} + ac\mathbb{E}_t x_{t+1} + cx_t \quad (6.59)$$

$$y_{t+2} = a\mathbb{E}_{t+2} y_{t+3} + cx_{t+2} \quad (6.60)$$

$$\implies \mathbb{E}_t y_{t+2} = a\mathbb{E}_t y_{t+3} + c\mathbb{E}_t x_{t+2} \quad (6.61)$$

$$\implies y_t = a^3 \mathbb{E}_t y_{t+3} + a^2 c \mathbb{E}_t x_{t+2} + ac\mathbb{E}_t x_{t+1} + cx_t \quad (6.62)$$

⋮

$$y_t = c \sum_{i=0}^T a^i \mathbb{E}_t x_{t+i} + a^{T+1} \mathbb{E}_t y_{t+T+1} \quad (6.63)$$

Let $T \rightarrow \infty$ and assume that

$$\lim_{T \rightarrow \infty} [a^{T+1} \mathbb{E}_t y_{t+T+1}] = 0 \quad (6.64)$$

i.e. we are ruling out bubble solution.

Then we get the fundamental solution:

$$y_t = c \sum_{i=0}^{\infty} a^i \mathbb{E}_t x_{t+i} \quad (6.65)$$

6.5 Stochastic Dynamic Discrete Time Infinite Horizon Model

$$\max \mathbb{E} \left[\sum_{t=1}^{\infty} \beta^t u(c_t) \right] \quad (6.66)$$

$$\text{s.t. } c_t + B_t = (1 + \rho_{t-1})B_{t-1} + y_t \quad (6.67)$$

Random Lagrangian Method (Kushner (1965))

FONC

$$\mathbb{E}_t [\beta^t u'(c_t) - \lambda_t] = 0 \quad (6.68)$$

$$\mathbb{E}_t [-\lambda_t + \lambda_{t+1}(1 + \rho_t)] = 0 \quad (6.69)$$

Note: At time t , variables dated t and earlier are known and hence they are not R.V.'s (random variables), i.e.

$$\mathbb{E}_t x_t = \mathbb{E}[x_t | I_t] = x_t \quad (6.70)$$

\implies

$$\beta^t u'(c_t) - \lambda_t = 0 \quad (6.71)$$

$$-\lambda_t + (1 + \rho_t) \mathbb{E}_t \lambda_{t+1} = 0 \quad (6.72)$$

\implies

$$u'(c_t) = (1 + \rho_t) \beta \mathbb{E}_t u'(c_{t+1}) \quad (6.73)$$

$$\mathbb{E}_t u'(c_{t+1}) = \frac{1}{(1 + \rho_t) \beta} u'(c_t) \quad (6.74)$$

$$u'(c_{t+1}) = \frac{1}{(1 + \rho_t)\beta} u'(c_t) + \epsilon_{t+1} \quad (6.75)$$

with $\mathbb{E}_t \epsilon_{t+1} = 0$. Hence it's almost an AR(1) process.

Claim: If $\mathbb{E}_t(x_{t+1}) = x_t$, then $x_{t+1} = x_t + \epsilon_{t+1}$, $E_t \epsilon_{t+1} = 0$.

Comment: This is also a regression equation/model that can be used for estimation, taken to data for test directly. If $u(c) = \ln c$, then

$$\frac{1}{c_{t+1}} = \frac{1}{(1 + \rho_t)\beta} \frac{1}{c_t} + \epsilon_{t+1} \quad (6.76)$$

which is a regression equation with time-varying coefficient (Kalman filter). (F. Mishkin 1986, NBER-U.C. Press)

Assumption: $\rho_t = \rho$ and $\beta = \frac{1}{1+\rho}$

Then

$$u'(c_{t+1}) = u'(c_t) + \epsilon_{t+1} \quad (6.77)$$

MU of consumption is a random walk.

With quadratic utility function, $MU = \alpha - \beta c$, Hall (1978) showed that

$$c_{t+1} = c_t + \epsilon_{t+1}. \quad (6.78)$$

References

[Sargent 1979] TOM SARGENT, “Chapter 9” *Macroeconomic Theory*, 1979